

Marginal and joint space representations within ABC, and the issue of bias

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Motivation

1. Specification of actual (marginal/joint) ABC target distribution did not explicitly appear within the literature
2. Must be possible to pose all ABC-SMC algorithms within a common framework (i.e. SMC samplers framework)
3. Biased/unbiased argument not resolved (within literature)
4. Algorithms with computational cost lower than $O(N^2)$ and whose particle weights don't degenerate to zero

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Background: Notation and Terminology

- Bayesian inference proceeds through the posterior distribution:

$$\pi(\theta|y) \propto \pi(y|\theta)\pi(\theta) \quad \theta \in \Theta, y \in \mathcal{X}$$

- Increasing interest in scenarios where the likelihood function is intractable
- Standard simulation procedures based on repeated likelihood evaluations (eg MCMC, SMC) require modification
- Collection of methods utilised for posterior simulation in the presence of intractable likelihood has become known as “Approximate Bayesian Computation” (ABC)

Joint Space Representation

- ABC methods employ standard procedure of embedding target posterior within augmented model from which sampling is viable

$$\pi(\theta, x|y) \propto \pi(y|x, \theta)\pi(x|\theta)\pi(\theta),$$

where $x \in \mathcal{X}$

- In ABC setting $x \sim \pi(x|\theta)$ is an artificial dataset distributed according to the likelihood
- $\pi(y|x, \theta)$ weights the intractable posterior, with high density in regions where summary statistics $T(x) \approx T(y)$ of x and y are similar

Weighting function

- Commonly utilised but “inefficient” form is:

$$\pi(y|x, \theta) \propto \begin{cases} 1 & \text{if } \rho(T(y), T(x)) \leq \epsilon \\ 0 & \text{else} \end{cases}$$

- Exclusively rewards those $T(x)$ within an ϵ -tolerance of $T(y)$ as measured by a distance function ρ
- Alternative forms of $\pi(y|x, \theta)$ have been evaluated, such as Gaussian and Epanechnikov kernels

Marginal Space Representation

- The target marginal distribution is:

$$\pi_{ABC}(\theta|y) \propto \int_{\mathcal{X}} \pi(y|x, \theta) \pi(x|\theta) \pi(\theta) dx$$

- If $\pi(y|x, \theta)$ for $T(y)=T(x)$ and 0 elsewhere, and if T is (“exactly”) sufficient, then $\pi_{ABC}(\theta|y) \equiv \pi(\theta|y)$ reduces to the true posterior

Marginal Space Representation: Monte Carlo Approximation

- The target marginal distribution can be approximated as follows:

$$\begin{aligned}\pi_{ABC}(\theta|y) &\propto \pi(\theta) \int_{\mathcal{X}} \pi(y|x, \theta) \pi(x|\theta) dx \\ &= \pi(\theta) \mathbb{E}_{\pi(x|\theta)} [\pi(y|x, \theta)] \\ &\approx \frac{\pi(\theta)}{S} \sum_{s=1}^S \pi(y|x^s, \theta),\end{aligned}$$

where $x^1, \dots, x^S \sim \pi(x|\theta)$ are draws from the intractable likelihood

ABC-MCMC: Targeting the marginal space

- Sisson et al (2008) argue that ABC-MCMC (Marjoram et al (2003)) is explicitly targeting the marginal distribution $\pi_{ABC}(\theta|y)$
- Proof of detailed balance in Marjoram et al (2003) relates to marginal distribution $\pi_{ABC}(\theta|y)$

ABC-MCMC Algorithm ($S = 1$)

At stage t :

1. Generate $\theta \sim q(\theta_t, \theta)$ from a proposal distribution.
2. Generate $x \sim \pi(x|\theta)$ from the likelihood.
3. With probability $\min \left\{ 1, \frac{\pi(y|x, \theta)\pi(\theta)q(\theta, \theta_t)}{\pi(y|x_t, \theta_t)\pi(\theta_t)q(\theta_t, \theta)} \right\}$ accept $\theta_{t+1} = \theta$, ($x_{t+1} = x$)
otherwise set $\theta_{t+1} = \theta_t$, ($x_{t+1} = x_t$).
4. Increment $t = t + 1$ and go to 1.

ABC-MCMC: Targeting the joint space

- The ABC-MCMC algorithm proceeds exactly as before

- MH-ratio:

$$\frac{\pi(\theta_{t+1}, x_{t+1}|y) q(\theta_t, x_t|\theta_{t+1}, x_{t+1})}{\pi(\theta_t, x_t|y) q(\theta_{t+1}, x_{t+1}|\theta_t, x_t)}$$

- Let $q(\theta_{t+1}, x_{t+1}|\theta_t, x_t) \triangleq \pi(x_{t+1}|\theta_{t+1})q(\theta_{t+1}|\theta_t)$

Then MH-ratio:

$$\frac{\pi(\theta_{t+1})\pi(y|x_{t+1}\theta_{t+1}) q(\theta_t|\theta_{t+1})}{\pi(\theta_t)\pi(y|x_t, \theta_t) q(\theta_{t+1}|\theta_t)}$$

- This is the same expression as in the marginal ABC-MCMC algorithm (if $S=I$), with the same sequence of algorithmic operations!

SMC Samplers: Notation

- SMC samplers operates on an artificial sequence of target distributions
- Defined to admit the desired target distribution as a marginal
- (IS) weight expression:

$$w_n(\theta_{1:n}) = \frac{\tilde{\phi}_n(\theta_{1:n})}{\tilde{\mu}_n(\theta_{1:n})} = \frac{\phi_n(\theta_n) \prod_{k=1}^{n-1} L_k(\theta_{k+1}, \theta_k)}{\mu_1(\theta_1) \prod_{k=2}^n M_k(\theta_{k-1}, \theta_k)}.$$

$$w_n(\theta_{n-1}, \theta_n) = \frac{\phi_n(\theta_n) L_{n-1}(\theta_n, \theta_{n-1})}{\phi_{n-1}(\theta_{n-1}) M_n(\theta_{n-1}, \theta_n)}.$$

ABC-SMC: targeting the marginal space

- Sequence of target distributions:

$$\phi_n(\theta_n) = \pi_{ABC,n}(\theta_n|y) \propto \int_{\mathcal{X}} \pi_n(y|x, \theta) \pi(x|\theta) \pi(\theta) dx$$

for $\epsilon_1 \geq \epsilon_2 \geq \dots \geq \epsilon_n$.

- E.g.

$$\pi_n(y|x, \theta) \propto \begin{cases} 1 & \text{if } \rho(T(y), T(x)) \leq \epsilon_n \\ 0 & \text{else,} \end{cases}$$

ABC-SMC: targeting the marginal space

- Weighting function (Sisson et al, 2007):

$$\begin{aligned} w_t &\propto \frac{\pi_{ABC,n}(\theta_n|y)}{\pi_{ABC,n}(\theta_{n-1}|y)} \frac{L_{n-1}(\theta_{n-1}|\theta_n)}{K_n(\theta_n|\theta_n)} \\ &\approx \frac{\frac{\pi(\theta_n)}{B_n} \sum_{b=1}^{B_n} \pi_n(y|x_n^b, \theta_n)}{\frac{\pi(\theta_{n-1})}{B_{n-1}} \sum_{b=1}^{B_{n-1}} \pi_{n-1}(y|x_{n-1}^b, \theta_{n-1})} \frac{L_{n-1}(\theta_{n-1}|\theta_n)}{K_n(\theta_n|\theta_n)} \end{aligned}$$

where this expression follows by inserting the Monte Carlo approximation to the marginal distribution into the numerator and denominator

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- With $B_N = 1$ expression can be compared to (marginal) MCMC

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- With $B_N = 1$ expression can be compared to (marginal) MCMC
- *Biased algorithm: ratio of Monte Carlo estimates*
 - Bias reduces for large B_N but can be significant for $B_N = 1$

ABC-SMC: targeting the marginal space

- Sisson et al, 2009 (errata note for Sisson et al, 2007 paper)
- Directly evaluate the importance sampling distribution, hence avoiding the need for Monte Carlo approximation in the denominator of $\pi_{ABC,n}(\theta_n|y)$
- This corresponds to using (an approximation to) the optimal L kernel

$$\begin{aligned}w_n(\theta_n) &= \frac{\pi_{ABC,n}(\theta_n|y)}{\int_{\Theta} \pi_{n-1}(\theta_{n-1}|y)K(\theta_n|\theta_{n-1})d\theta_{n-1}} \\ &\approx \frac{\pi(\theta_n)}{\sum_{j=1}^N w_{n-1}(\theta_{n-1}^{(j)})k(\theta_n|\theta_{n-1})}\end{aligned}$$

- Unbiased algorithm, but now with $O(N^2)$ computation

ABC-SMC: targeting the joint space

- Incremental weight on joint space given by:

$$w_n[(\theta_{n-1}, x_{n-1}), (\theta_n, x_n)] \propto \frac{\pi_n(\theta_n, x_n | y) L_{n-1}[(\theta_n, x_n), (\theta_{n-1}, x_{n-1})]}{\pi_{n-1}(\theta_{n-1}, x_{n-1} | y) M_n[(\theta_{n-1}, x_{n-1}), (\theta_n, x_n)]}$$

- As in ABC-MCMC, interpretation on joint space is through the use of “standard” SMC sampler methodology
- For general choice of L_{n-1} , a (generally intractable) normalising constant appears in weight expression which renders this model impractical

Some unbiased SMC-ABC algorithms

- ABC-SMC(MCMC)

$$w_n[(\theta_{n-1}, x_{n-1}), (\theta_n, x_n)] \propto \frac{\pi_n(\theta_{n-1}, x_{n-1} | y)}{\pi_{n-1}(\theta_{n-1}, x_{n-1} | y)} = \frac{\pi_n(y | x_{n-1}, \theta_{n-1})}{\pi_{n-1}(y | x_{n-1}, \theta_{n-1})},$$

- Note that it is possible for all weights to be assigned zero weight
 - Del Moral et al (2008) introduce automated ϵ_n selection scheme
 - Sisson et al (2007) include a PRC step
- ABC-SMC(PMC)
 - $O(N^2)$ algorithm

Application

- Estimate: number and parameters of a CBRN release
- Variable number of releases of a given material, specified by:
 - Release location; Release times; Release amounts
- Grid of sensors located downwind of the releases
- Dispersion model for how the plumes propagate over time
 - This can be a complex stochastic process, varying dependent on atmospheric conditions (e.g. wind profiles), and the environment (e.g. buildings)
- Sensor model which relates the actual concentration to the measurements returned by the sensors
 - Again can be complex

ABC-SMC (RJMCMC)

- Use of Trans-dimensional Metropolis-Hastings kernel within ABC-SMC
- As usual, difficult to specify efficient transition distribution
 - Results in severe particle depletion
- Need to resolve this issue, or look at alternative methodology

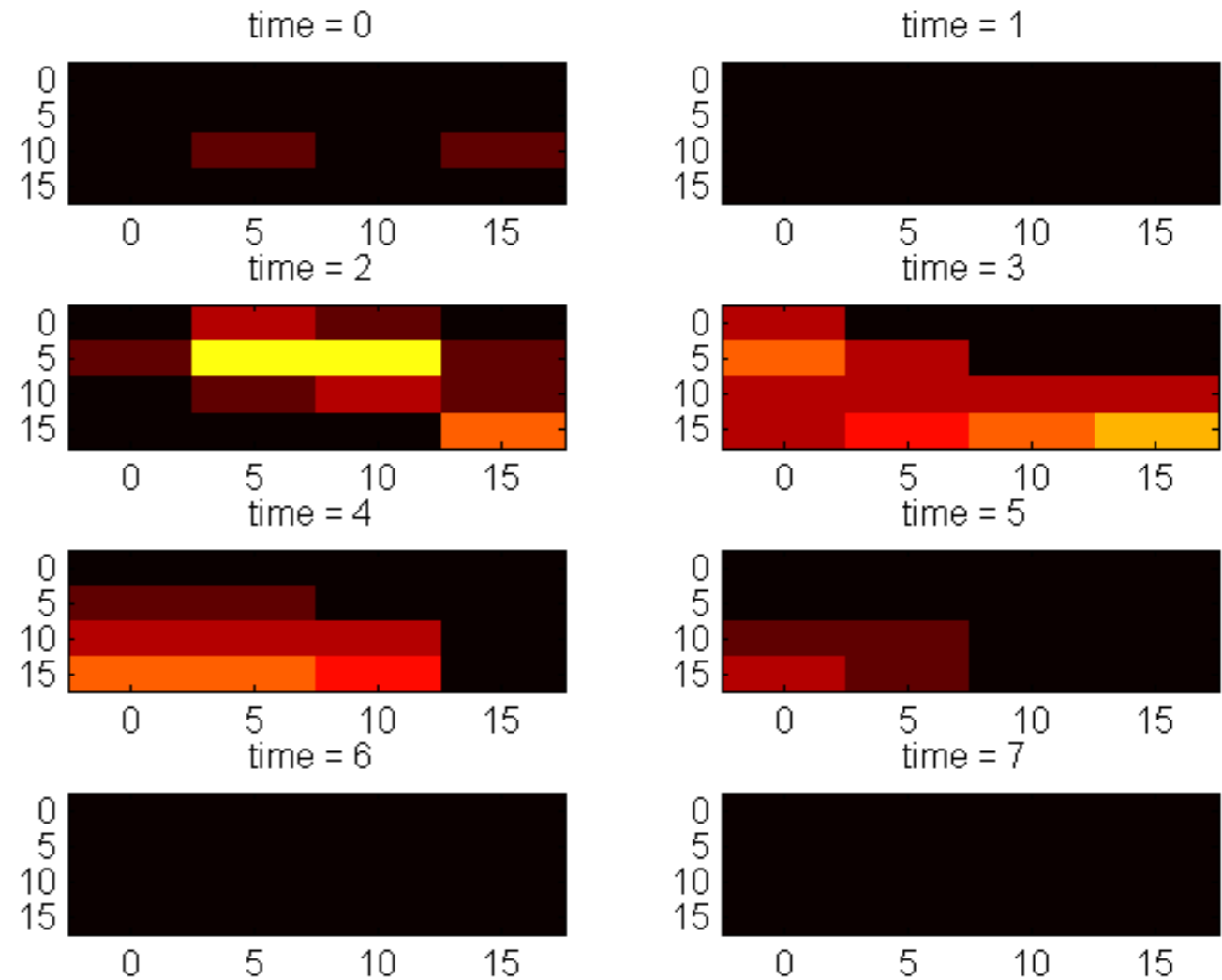
Three releases

- Simulated three-release scenario

Release	Time	x-coord	Mass
1 st	0	(40,15)	250
2 nd	0	(30,5)	300
3 rd	0.8	(40,15)	200

Ground-truth release details

Cross wind of (-11,0)



Sensor measurements (4-by-4 arrays)

Results

Release	MAP number of releases	Time	x-coord	y-coord	Mass
<i>Ground-truth</i>	3	0 0 0.8	40 30 40	15 5 15	250 300 200
ABC-SMC	Fixed = 3	-0.43 -0.21 0.1	45.0 32.0 45.9	15.7 5.2 16.0	205.7 263.0 198.7
Trans-dimensional ABC-SMC	3	-0.97 -0.03 -0.36	53.4 30.4 49.6	16.7 5.3 15.2	302.9 261.6 162.5

- Although errors exist in the values, once effects from early or late release predictions are taken into account one can clearly link the estimates to the ground-truth values.
- Hence one can obtain reasonable estimates, even when the number of releases is unknown.

Conclusions

1. *Specification of actual (marginal/joint) ABC target distribution did not explicitly appear within the literature*
Marginal and joint space representations exposed (explicitly)
 2. *Must be possible to pose all ABC-SMC algorithms within a common framework (i.e. SMC samplers framework)*
Clarity on use of ABC-SMC for both marginal and joint space representations
 3. *Biased/unbiased argument not resolved (within literature)*
Bias argument now (hopefully) resolved
 4. *Algorithms with computational cost lower than $O(N^2)$ and whose particle weights don't degenerate to zero*
Only one SMC algorithm that has cost lower than $O(N^2)$?
- Novel application of ABC-SMC(RJ-MCMC) presented

- Questions?