




ABC SMC for dynamical systems

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ABC in Paris, 26/06/2009



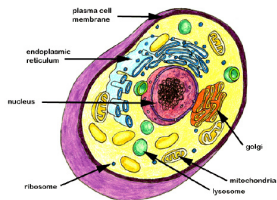
Outline

- 1 ABC for dynamical systems
- 2 ABC based on Sequential Monte Carlo
- 3 ABC SMC for model selection
- 4 Applications
- 5 Perturbation kernels

Motivation

Complex biological systems

- Models often ODE or stochastic master equations
- High dimensional parameter space
- Time course, non-equidistant, missing data
- Complex likelihood surface



Interested in

- Characterization of distributions over parameters rather than point estimates
- Parameter sensitivity
- Hypothesis testing: nested / non nested models
- Computational efficiency

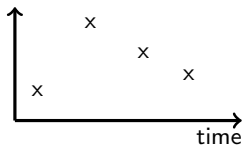
ABC framework for dynamical systems



Prior
 $P(\theta)$



Posterior
 $P(\theta|D)$



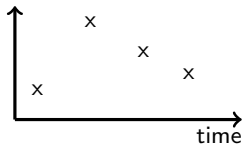
ABC framework for dynamical systems



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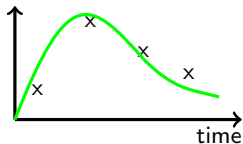
ABC framework for dynamical systems



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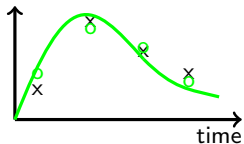
ABC framework for dynamical systems



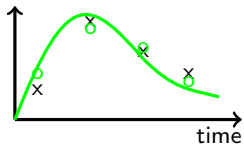
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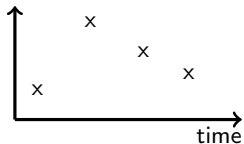
Posterior
 $P(\theta|D)$



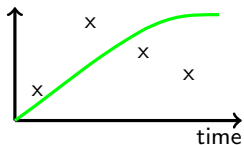
ABC framework for dynamical systems

Prior
 $P(\theta)$ Posterior
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ABC framework for dynamical systems

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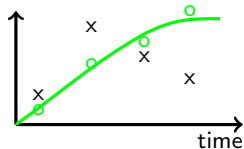
ABC framework for dynamical systems



Prior
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Posterior
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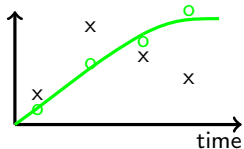
ABC framework for dynamical systems



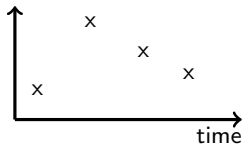
Prior
 $P(\theta)$



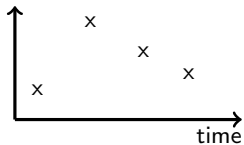
Posterior
 $P(\theta|D)$



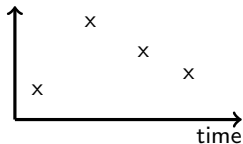
ABC framework for dynamical systems

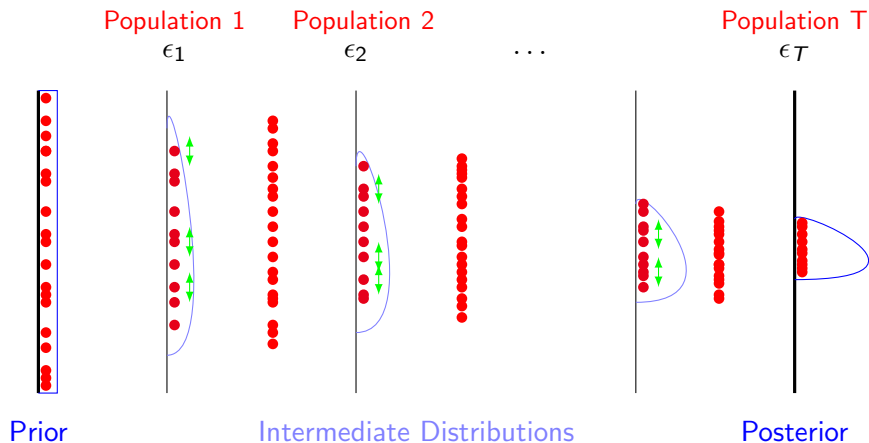
Prior
 $P(\theta)$ Posterior
 $P(\theta|D)$ 

ABC framework for dynamical systems

Prior
 $P(\theta)$ Posterior
 $P(\theta|D)$ 

ABC framework for dynamical systems

Prior
 $P(\theta)$ Posterior
 $P(\theta|D)$ 

ABC SMC for estimation of $P(\theta|D)$ 

(Sisson et al., 2007)

Weights

Sisson *et al.* (2007)

$$w_t^{(i)} = \frac{\pi(\theta_t^{(i)}) L_{t-1}(\theta^* | \theta_t^{(i)})}{\pi(\theta^*) K_t(\theta_t^{(i)} | \theta^*)} \propto 1 \quad \text{when } L_{t-1} = K_t$$

Instead:

$$w_t(\theta_t) = \frac{\pi_t(\theta_t)}{\eta_t(\theta_t)}$$

$$\eta_t(\theta_t) = 1(\pi(\theta_t) > 0) 1(\text{dist} > \epsilon_t) \int \pi_{t-1}(\theta_{t-1}) K_t(\theta_t | \theta_{t-1}) d\theta_{t-1}$$

$$w_t^{(i)} = \frac{\pi(\theta_t^{(i)})}{\sum_{j=1}^N w_{t-1}^{(j)} K_t(\theta_t^{(i)} | \theta_{t-1}^{(j)})}$$

Approximate Bayesian computation scheme for parameter inference and model selection in dynamical systems

Tina Toni^{1,2,*}, David Welch^{3,†}, Natalja Strelkova⁴,
Andreas Ipsen⁵ and Michael P. H. Stumpf^{1,2,*}

Related

- S. A. Sisson, Y. Fan and M. M. Tanaka. A note on backward kernel choice for sequential Monte Carlo without likelihoods. (S. Sisson's webpage)
- Beaumont MA, Cornuet JM, Marin JM and Robert CP. Adaptive approximate Bayesian computation. (arXiv)

Model selection: $P(M|D) = ?$

1. Marginal likelihoods

- For each model separately estimate $P(D|M)$:

$$P(D|M) = \int_{\theta} P(D|\theta, M)P(\theta|M)d\theta$$

- Then

$$P(M|D) = \frac{P(D|M)P(M)}{\sum_{M'} P(D|M')P(M')}.$$

2. Joint space

- Include model M as an extra parameter:
 $(M, \theta^{(1)}, \dots, \theta^{(M)})$
- ABC SMC
 $\rightarrow P(M, \theta^{(1)}, \dots, \theta^{(M)}|D)$
- Marginalize
 $\rightarrow P(M|D)$

Example: ABC rejection

1. Marginal likelihoods

- $M = 1$, N_1 particles:

$$P(D|M = 1) = \frac{\#\text{accepted}_1}{N_1}$$

- $M = 2$, N_2 particles:

$$P(D|M = 2) = \frac{\#\text{accepted}_2}{N_2}$$

- ...

$$P(M = i|D) =$$

$$\frac{\frac{\#\text{accepted}_i}{N_i} P(M = i)}{\sum_j \frac{\#\text{accepted}_j}{N_j} P(M = j)}$$

(Wilkinson, PhD Thesis)

2. Joint space

- Propose M .
- Propose $\theta^{(M)}$.
- Accept/reject $(M, \theta^{(M)})$.

$$P(M = i|D) =$$

$$\frac{\#\text{accepted particles with } M = i}{\#\text{all accepted particles}}$$

(Grelaud et al., ArXiv)

How can we use ideas from SMC in these approaches?

Model selection on a joint space

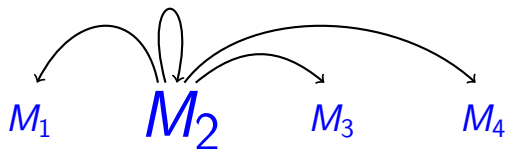
 M_1 M_2 M_3 M_4

Model selection on a joint space

M_1 M_2 M_3 M_4

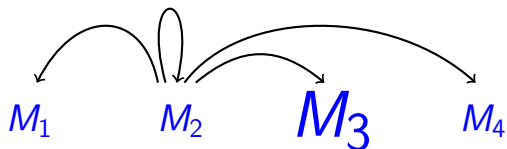
M^*

Model selection on a joint space



$$M^*$$
$$M^{**} \sim KM(M|M^*)$$

Model selection on a joint space

 M^*

$$M^{**} \sim KM(M|M^*)$$

Model selection on a joint space

(M_3, θ_3)
 (M_3, θ_7)
 (M_3, θ_6) (M_3, θ_2)
 (M_3, θ_8) (M_3, θ_5) (M_3, θ_1)
 (M_3, θ_4)
 (M_3, θ_9)

$$M^*$$

$$M^{**} \sim KM(M|M^*)$$

$$\theta^*$$

Model selection on a joint space

(M_3, θ_3)
 (M_3, θ_7)
 (M_3, θ_6) (M_3, θ_2)
 (M_3, θ_8) **(M_3, θ_5)** (M_3, θ_1)
 (M_3, θ_4)
 (M_3, θ_9)

$$M^*$$

$$M^{**} \sim KM(M|M^*)$$

$$\theta^*$$

Model selection on a joint space

(M_3, θ_3)
 (M_3, θ_7)
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 (M_3, θ_4)
 (M_3, θ_9)

 M^*

$$M^{**} \sim KM(M|M^*)$$

 θ^*

$$\theta^{**} \sim KP(\theta|\theta^*)$$

Model selection on a joint space

$$(M^{**}, \theta^{**})$$

$$M^*$$

$$M^{**} \sim KM(M|M^*)$$

$$\theta^*$$

$$\theta^{**} \sim KP(\theta|\theta^*)$$

accept / reject

Model selection on a joint space

$$w(M^{**}, \theta^{**})$$

 M^*

$$M^{**} \sim KM(M|M^*)$$

 θ^*

$$\theta^{**} \sim KP(\theta|\theta^*)$$

accept / reject

calculate w

Model selection on joint space: Weight calculation

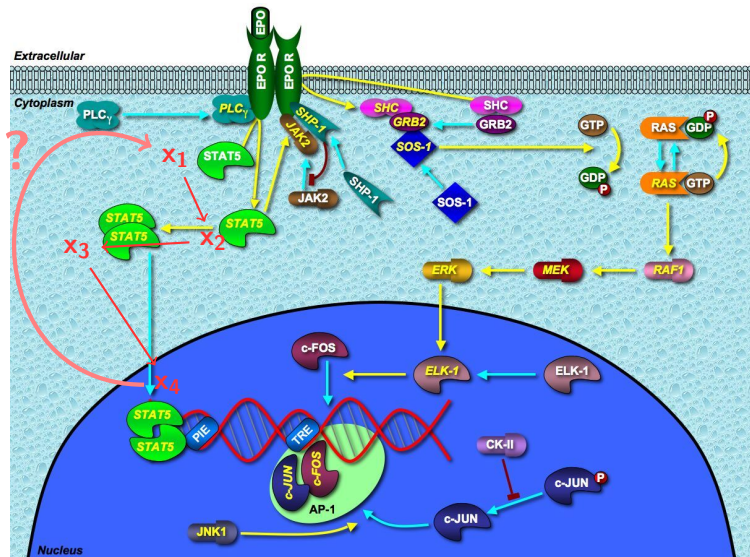
$$w_t(m^{**}, \theta^{**})$$

$$=$$

$$\pi(m^{**}, \theta^{**})$$

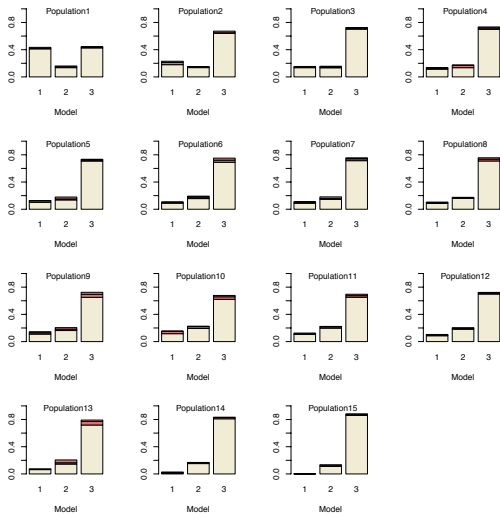
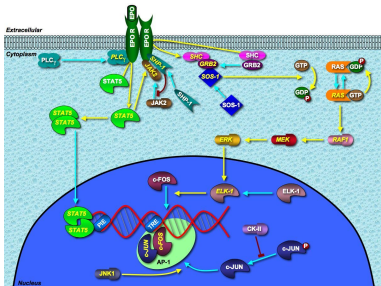
$$\underbrace{\sum_{j=1}^{\mathcal{M}} P_{t-1}^{(j)} KM_t(m^{**} | m_{t-1}^{(j)})}_{\text{model perturbation}} \underbrace{\sum_{k; m_{t-1}=m^{**}} \frac{w_{t-1}^{(k)}}{\sum_{l; m_{t-1}=m^{**}} w_{t-1}^{(l)}} KP_{t, m^{**}}(\theta^{**} | \theta_{t-1}^{(k)})}_{\text{parameter perturbation}}$$

Model selection: JAK-STAT signalling pathway



(adapted from Biocarta)

Model selection: JAK-STAT signalling pathway



Model selection: Phosphorylation dynamics

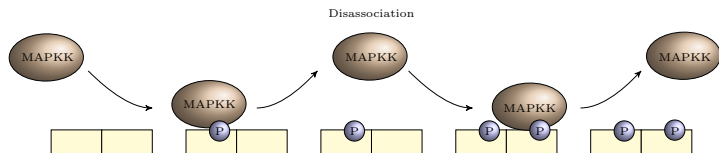


Figure 1: Distributive phosphorylation.

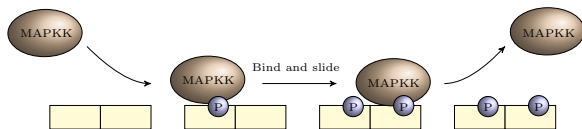
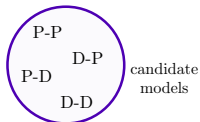
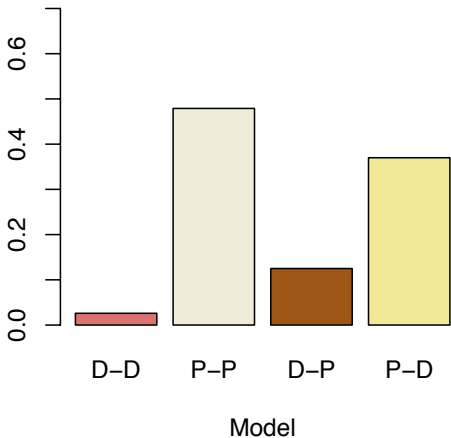


Figure 2: Processive phosphorylation.

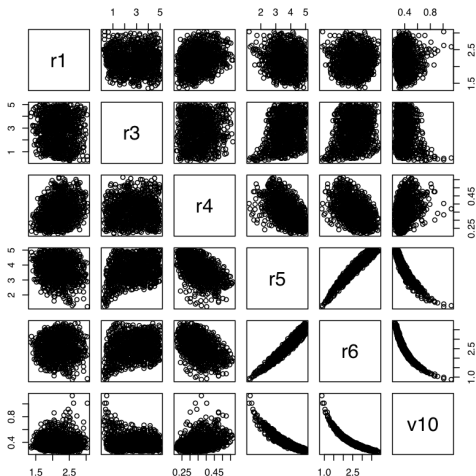
Model selection: Phosphorylation dynamics



Marginal posterior distribution P(MID)



Parameter estimation: Sensitivity, Identifiability



Pros and Cons

Advantages

- Missing data, non-equidistant time intervals
- Computational improvement over ABC REJ
- Parameter sensitivity for free

Disadvantages

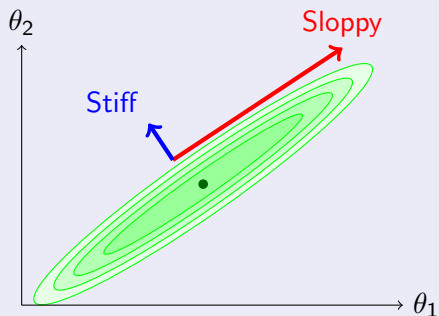
- Computational cost \rightarrow parallelization
- User input: tolerance schedule \rightarrow Doucet & Jasra
- Sensitivity to changes in $\epsilon_T \rightarrow$ regression adjustment Leuenberger ($11^{00} - 11^{30}$), Blum ($15^{00} - 15^{30}$), Francois ($17^{00} - 17^{30}$)

Perturbation kernels

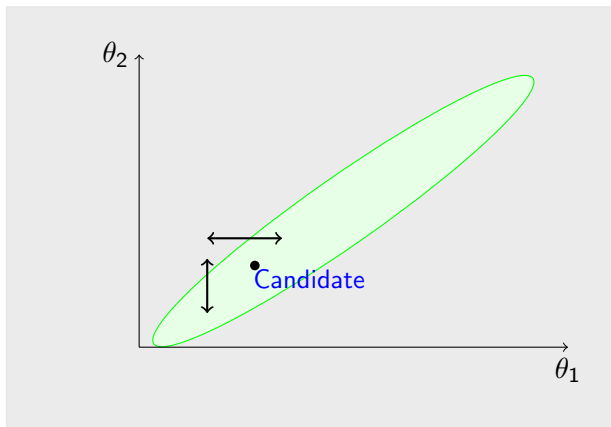
Component-wise

- Gaussian
- Uniform

Parameters in systems biology

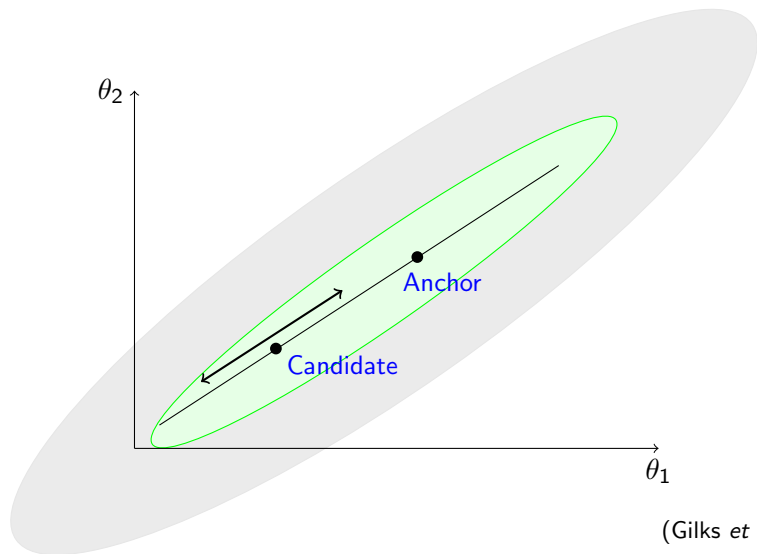


Component-wise perturbation

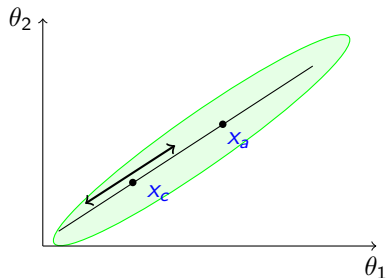


Idea: Perturbation kernels from evolutionary computation and snookering.

Snooker perturbation



Snooker perturbation



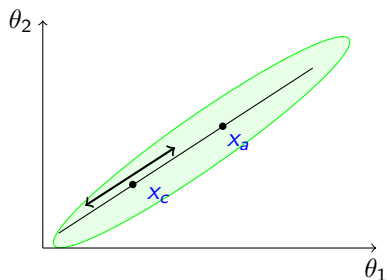
$$x'_c = x_c + r \frac{x_a - x_c}{\|x_a - x_c\|}$$

$$r \sim N(0, \|x_a - x_c\|)$$

$$r \sim U(-1, 1) \|x_a - x_c\|$$

$$w_t(x'_c) = \frac{\pi(x'_c)}{\sum \frac{w_{x_c} w_{x_a}}{1 - w_{x_c}} K_t(x'_c | x_a, x_c)}$$

Snooker perturbation



$$x'_c = x_c + r \frac{x_a - x_c}{\|x_a - x_c\|}$$

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Challenge: How to avoid reduced particle diversity?

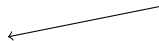
Crossover

$$(a_1 \ a_2 \ a_3 \mid a_4 \ a_5)$$

$$(b_1 \ b_2 \ b_3 \mid b_4 \ b_5)$$

$$(a_1 \ a_2 \ a_3 \ b_4 \ b_5)$$

← Parent chromosomes



Candidate particle

Crossover

$$\begin{array}{l} (a_1 \ a_2 \ a_3 \mid a_4 \ a_5) = x_{t-1}^{(i)} \\ (b_1 \ b_2 \ b_3 \mid b_4 \ b_5) = x_{t-1}^{(j)} \end{array} \quad \begin{array}{l} \longleftarrow \\ \longleftarrow \end{array} \quad \text{Parent chromosomes}$$

$$(a_1 \ a_2 \ a_3 \ b_4 \ b_5) = x_t \quad \text{Candidate particle}$$

$$w(x_t) = \pi(x_t) / \left(\underbrace{\sum_i \left[(1 - \rho) w_{t-1}^{(i)} K_t(x_t | x_{t-1}^{(i)}) \right]}_{\text{perturbation}} + \underbrace{\rho \frac{1}{N(\rho - 1)} \sum_j w_{t-1}^{(i)} w_{t-1}^{(j)} \mathbb{1}(x_t \text{ is recombination of } x_{t-1}^{(i)}, x_{t-1}^{(j)})}_{\text{recombination}} \right)$$

ρ : probability
of recombination

ρ : particle
length

Summary

- ABC SMC for parameter estimation and model selection of dynamical systems.
- Challenges connected to posterior parameter distributions of biological systems:
 - Bayesian framework suited to study parameter sensitivity
 - Sloppiness
 - Perturbation kernels should exploit our knowledge about posterior distributions

Acknowledgements

- David Welch (Penn State University)
- Mark Beaumont (University of Reading)
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- David Balding, Ajay Jasra (Imperial College London)
- Paul Kirk