

**Exercise sheet n°1 : Poisson processes**

**Exercise 1.** Let  $(\tau_n, n \geq 1)$  be a sequence of IID (independent and identically distributed) non-negative random variables. Set  $T_0 = 0$ ,

$$T_n = \tau_1 + \dots + \tau_n, \quad n \geq 1,$$

and

$$N_t = \#\{i \geq 1 : T_i \leq t\}, \quad t \geq 0.$$

1. Give a necessary and sufficient condition for having

$$\mathbb{P}(N \text{ only makes jumps of size 1}) > 0.$$

2. Show that under this condition

$$\mathbb{P}(N \text{ only makes jumps of size 1}) = 1.$$

3. Is it possible that  $(T_n, n \geq 0)$  converges to a finite limit with positive<sup>1</sup> probability ?
4. Compute the probability of the event  $\{\exists t \geq 0 : N(t) = \infty\}$ .

**Exercise 2.** Let  $N$  be a Poisson process with intensity  $\lambda > 0$ . Prove and give an interpretation of the following properties

1.  $\mathbb{P}(N_h = 1) = \lambda h + o(h)$  ( $h \rightarrow 0$ )
2.  $\mathbb{P}(N_h \geq 2) = o(h)$  ( $h \rightarrow 0$ )
3.  $\mathbb{P}(N_h = 0) = 1 - \lambda h + o(h)$  ( $h \rightarrow 0$ ).
4.  $\forall t \geq 0, \mathbb{P}(N \text{ jumps at time } t) = 0$ .
5. Compute  $\text{Cov}(N_s, N_t), \forall s, t \geq 0$ .

**Exercise 3.** Let  $N$  be a counting process with stationary and independent increments. Assume that there exists  $\lambda > 0$  such that

$$\mathbb{P}(N_h = 1) = \lambda h + o(h), \quad \mathbb{P}(N_h \geq 2) = o(h).$$

For  $u \in \mathbb{R}$ , let  $g_t(u) = \mathbb{E}[e^{iuN_t}]$ .

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1. En anglais le mot *positive* signifie strictement positif. Pour dire positif au sens large on dit *non-negative*. De même les termes *negative*, *bigger*, *smaller* sont à prendre au sens strict.

1. Prove that  $g_{t+h}(u) = g_t(u)g_h(u)$  for every  $t, h \geq 0$ .

2. Prove that

$$\frac{d}{dt}g_t(u) = \lambda(e^{iu} - 1)g_t, \quad g_0(u) = 1.$$

3. Conclude.

**Exercise 4.** Let  $N$  be a Poisson process with intensity  $\lambda > 0$ , modelling the arrival times of the claims for an insurance company. Let  $T_1$  denote the arrival time of the first claim. Show that the conditional law of  $T_1$  given  $N_t = 1$  is uniformly distributed over  $[0, t]$ .

**Exercise 5.** Let  $(T_n, n \geq 0)$  ( $T_0 = 0$ ) be a renewal process and  $N$  its associated counting process. Assume that  $N$  has independent and stationary increments.

1. Show that

$$\mathbb{P}(T_1 > s + t) = \mathbb{P}(T_1 > t)\mathbb{P}(T_1 > s), \quad \forall s, t \geq 0.$$

2. Derive that  $N$  is a Poisson process.

**Exercise 6.**

1. Show that two independent Poisson processes cannot jump simultaneously a.s.

2. Let  $N^1$  and  $N^2$  be two independent Poisson processes with parameters  $\lambda_1 > 0$  and  $\lambda_2$  respectively. Show that the process

$$N_t = N_t^1 + N_t^2, \quad t \geq 0$$

is a Poisson process and give its intensity.

3. Derive that the sum of  $n$  independent Poisson processes with respective intensities  $\lambda_1 > 0, \dots, \lambda_n > 0$  is a Poisson process and give its intensity.

**Exercise 7.** Liver transplants arrive at a hospital according to a Poisson process  $N$  with intensity  $\lambda > 0$ . Two patients are waiting for a transplant. The first patient has lifetime  $T$  (before the transplant) according to an exponential distribution with parameter  $\mu_1$ . The second one has lifetime  $T'$  (before the transplant) according to an exponential distribution with parameter  $\mu_2$ . The rule is that the first transplant that arrives at the hospital is given to the first patient if he/she is still alive, and to the second patient otherwise. Assume that  $T, T'$  and  $N$  are independent.

1. Compute the probability that the first patient is transplanted.

2. Compute the probability that the second patient is transplanted.

3. Let  $X$  denote the number of transplants arrived at the hospital during  $[0, T]$ . Compute the law of  $X$ .

**Exercise 8.** [The bus paradox] Buses arrive at a given bus stop according to a Poisson process with intensity  $\lambda > 0$ . You arrive at the bus stop at time  $t$ .

1. Give a first guess for the value of the average waiting time before the following bus arrives ?
2. Let  $B_t = T_{N_t+1} - t$  be the waiting time before the next bus, and let  $A_t = t - T_{N_t}$  denote the elapsed time since the last bus arrival. Compute the joint distribution of  $(A_t, B_t)$  (hint : compute first  $\mathbb{P}(A_t \geq x_1, B_t \geq x_2)$  for  $x_1, x_2 \geq 0$ ).
3. Derive that the random variables  $A_t$  and  $B_t$  are independent. What are their distributions ?
4. In particular, compute  $\mathbb{E}[B_t]$ . Compare with your initial first guess.

**Exercise 9.** [Law of large numbers and central limit theorem.]

1. Recall and prove a law of large numbers for a Poisson process with intensity  $\lambda > 0$ .
2. Prove that  $N$  satisfies the following central limit theorem

$$\frac{N_t - \lambda t}{\sqrt{\lambda t}} \xrightarrow{\text{law}} \mathcal{N}(0, 1) \text{ as } t \rightarrow \infty,$$

- (a) by using characteristic functions
- (b) by showing first that  $(N_n - \lambda n) / \sqrt{\lambda n}$  converges in distribution as  $n \rightarrow \infty$  and then  $\max_{t \in [n, n+1]} (N_t - N_n) / \sqrt{n} \rightarrow 0$  in probability.

**Exercise 10.**

1. Give an expression for the density function of the conditional distribution of

$$(T_1, \dots, T_n) \text{ given } N_t = n$$

when  $N$  is a Poisson process with intensity  $\lambda$  and  $0 < T_1 < \dots < T_n < \dots$  are its jump times.

2. Derive an expression for the density of  $T_i$  given  $N_t = n$ ,  $\forall 1 \leq i \leq n$  and similarly for  $(T_i, T_j)$  given  $N_t = n$ ,  $\forall 1 \leq i < j \leq n$ .
3. Set  $U_{i,j} = T_j - T_i$ ,  $1 \leq i < j \leq n$ . Give an expression for the density of  $U_{i,j}$  given  $N_t = n$ . Derive an expression for the density of  $T_n - T_{n-1}$  given  $N_t = n$ .

**Exercise 11.** Let  $X = (X_t, t \geq 0)$  be a continuous time process and  $\mathcal{F} = (\mathcal{F}_t, t \geq 0)$  a *filtration*, i.e. a nested family of sigma-fields  $\mathcal{F}_s \subset \mathcal{F}_t \subset \mathcal{A} \ \forall s \leq t$ , where  $\mathcal{A}$  is the sigma-field on the probability space  $(\Omega, \mathcal{A}, \mathbb{P})$  over which  $X$  is defined. The process  $X$  is a *martingale* with respect to the filtration  $\mathcal{F}$  if  $X_t$  is  $\mathcal{F}_t$ -measurable and integrable  $\forall t$  and

$$\mathbb{E}[X_t | \mathcal{F}_s] = X_s, \quad \forall 0 \leq s \leq t.$$

Let  $N = (N_t, t \geq 0)$  be a Poisson process with intensity  $\lambda > 0$ . Show that the three processes

1.  $(N_t - \lambda t, t \geq 0);$
2.  $((N_t - \lambda t)^2 - \lambda t), t \geq 0);$
3.  $(\exp(uN_t + \lambda t(1 - e^u)), t \geq 0)$  (for a given real number  $u$ );

are martingales with respect to the filtration generated by  $N$ , i.e.  $\mathcal{F}_t^N = \sigma(N_s, s \leq t)$ .

**Exercise 12.** Let  $N$  be a Poisson process with intensity  $\lambda > 0$  and let  $0 < T_1 < \dots < T_n < \dots$  denote its jump times.

1. Show that  $T_n/n$  converges almost surely as  $n \rightarrow \infty$  and identify its limit.
2. Show that  $\sum_{i \geq 1} T_i^{-2}$  converges almost surely. Let  $X$  denote its limit.
3. Show that  $X_{N_t} = \sum_{i=1}^{N(t)} T_i^{-2} \rightarrow X$  a.s. as  $t \rightarrow \infty$ .
4. Let  $(U_i, i \geq 1)$  denote a sequence of independent uniform random variables on  $[0, 1]$ . We admit the following result

$$n^{-2} \sum_{i=1}^n U_i^{-2} \xrightarrow[n \rightarrow \infty]{\text{law}} Z,$$

where  $Z$  is a positive random variable, whose Laplace transform is given by  $\mathbb{E}[\exp(-sZ)] = \exp(-c\sqrt{s}), \forall s \geq 0$ , for some  $c > 0$ . **The goal is to show that  $X$  and  $c'Z$  have same law for some  $c'$  that we will explicitly compute.**

We assume moreover that  $(U_i, i \geq 1)$  is independent of  $N$ .

- (a) Show that for every  $n \geq 1$  and every  $t > 0$ , the law of  $X_{N_t}$  given  $N_t = n$  is the same as the law of  $t^{-2} \sum_{i=1}^n U_i^{-2}$ .
- (b) Derive that  $X_{N(t)}$  has same distribution as  $t^{-2} \sum_{i=1}^{N(t)} U_i^{-2}$ .
- (c) Prove that

$$N(t)^{-2} \sum_{i=1}^{N(t)} U_i^{-2} \xrightarrow{\text{law}} Z \quad \text{as } t \rightarrow \infty.$$

- (d) Recall the law of large numbers for Poisson processes and conclude.
5. Derive  $\mathbb{E}[X] = \infty$ .

**Exercise 13.** Let  $N = (N_t, t \geq 0)$  be a standard Poisson process with intensity  $\lambda > 0$ . Let  $f : [0, \infty) \rightarrow [0, \infty)$  be a locally bounded Borel function. Set

$$N(f)_t = \sum_{i \geq 1} f(T_i) \mathbf{1}_{\{T_i \leq t\}} \quad \text{for } t \geq 0,$$

where the  $(T_i)_{i \geq 1}$  are the jump times of  $N$ .

1. Show that for all  $t \geq 0$ , we have  $N(f)_t < \infty$  almost-surely.

2. If  $f(s) = \mathbf{1}_{(a,b]}(s)$  where  $[a, b] \subset [0, t]$ , what is the distribution of  $N(\mathbf{1}_{(a,b]})_t$  ?
3. Show that for  $u \geq 0$ , we have

$$\mathbb{E} [e^{-uN(f)_t} | N_t = n] = \frac{1}{t^n} \left( \int_0^t e^{-uf(s)} ds \right)^n.$$

4. Derive  $\mathbb{E} [e^{-uN(f)_t}]$  and find back the result of Question 2.
5. Compute  $\mathbb{E} [N(f)_t]$  and  $\text{Var}[N(f)_t]$ .
6. Prove that  $N(f)_t - \lambda \int_0^t f(s) ds$  is a martingale.